
Empirical Methods in CF

Lecture 2 – Causality

Professor Todd Gormley

Background readings for lecture

- Roberts-Whited
 - *Section 2*
- Angrist and Pischke
 - *Section 3.2*
- Wooldridge
 - *Sections 4.3 & 4.4*
- Greene
 - *Sections 5.8-5.9*

Motivation

- As researchers, we are interested in making causal statements
 - Ex. #1 – what is the *effect* of a change in corporate taxes on firms' leverage choice?
 - Ex. #2 – what is the *effect* of giving a CEO more stock ownership in the firm on the CEO's desire to take on risky investments?
- I.e. we don't like to just say variables are 'associated' or 'correlated' with each other

What do we mean by causality?

- Recall from earlier lecture, that if our linear model is the following...

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + u$$

And, we want to infer β_1 as the causal effect of x_1 on y , holding all else equal, then we need to make the following assumptions...

The basic assumptions

- *Assumption #1: $E(u) = 0$*
- *Assumption #2: $E(u | x_1, \dots, x_k) = E(u)$*
 - In words, average of u (i.e. unexplained portion of y) does not depend on value of x
 - This is “conditional mean independence” (CMI)
- Generally speaking, you need the estimation error to be uncorrelated with all the x 's

Unbiasedness *versus* Consistency

- When we say an estimate is unbiased or consistent, it means we think it has a causal interpretation...
 - I.e. the CMI assumption holds and the x 's are all uncorrelated with the disturbance, u
 - **Bias** refers to finite sample property; **consistency** refers to asymptotic property
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More formally...

- An estimate, $\hat{\beta}$, is unbiased if $E(\hat{\beta}) = \beta$
 - I.e. on average, the estimate is centered around the true, unobserved value of β
 - Doesn't say whether you get a more precise estimate as sample size increases
 - An estimate is consistent if $\underset{N \rightarrow \infty}{plim} \hat{\beta} = \beta$
 - I.e. as sample size increases, the estimate converges (in probability limit) to the true coefficient
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Tangent – CMI versus correlation

- CMI (which implies x and u are uncorrelated) is needed for unbiasedness [*which is again a finite sample property*]
- But, we only need to assume a zero correlation between x and u for consistency [*which is a large sample property*]
- **This is why I'll typically just refer to whether u and x are correlated in my test of whether we can make causal inferences**

Three main ways this will be violated

- Omitted variable bias
- Measurement error bias
- Simultaneity bias

- Now, let's go through each in turn...

Omitted variable bias (OVB)

- Probably the most common concern you will hear researchers worry about
- **Basic idea** = the estimation error, u , contains other variable, e.g. z , that affects y **and** is correlated with an x
 - **Please note!** The omitted variable is only problematic if correlated with an x

OVB more formally, with one variable

- You estimate: $y = \beta_0 + \beta_1 x + u$
- But, true model is: $y = \beta_0 + \beta_1 x + \beta_2 z + v$
- Then, $\hat{\beta}_1 = \beta_1 + \delta_{xz} \beta_2$, where δ_{xz} is the coefficient you'd get from regressing the omitted variable, z , on x ; and

$$\delta_{xz} = \frac{\text{cov}(x, z)}{\text{var}(x)}$$

Interpreting the OVB formula

$$\hat{\beta}_1 = \beta_1 + \frac{\text{cov}(x, z)}{\text{var}(x)} \beta_2$$

Effect of x on y Regression of z on x Effect of z on y

Bias

- Easy to see, estimated coefficient is only unbiased if $\text{cov}(x, z) = 0$ [i.e. x and z are uncorrelated] **or** z has no effect on y [i.e. $\beta_2 = 0$]

Direction and magnitude of the bias

$$\hat{\beta}_1 = \beta_1 + \frac{\text{cov}(x, z)}{\text{var}(x)} \beta_2$$

- Direction of bias given by signs of β_2 , $\text{cov}(x, z)$
 - E.g. If know z has positive effect on y [i.e. $\beta_2 > 0$] and x and z are positively correlated [$\text{cov}(x, z) > 0$], then the bias will be positive
- Magnitude of the bias will be given by magnitudes of β_2 , $\text{cov}(x, z) / \text{var}(x)$

Example – One variable case

- Suppose we estimate: $\ln(wage) = \beta_0 + \beta_1 educ + w$
- But, true model is:

$$\ln(wage) = \beta_0 + \beta_1 educ + \beta_2 ability + u$$

- What is likely bias on $\hat{\beta}_1$? Recall,

$$\hat{\beta}_1 = \beta_1 + \frac{\text{cov}(educ, ability)}{\text{var}(educ)} \beta_2$$

Example – Answer

- Ability & wages likely positively correlated, so $\beta_2 > 0$
- Ability & education likely positive correlated, so $\text{cov}(\text{educ}, \text{ability}) > 0$
- Thus, the bias is likely to positive! $\hat{\beta}_1$ is too big!

OVB – General Form

- Once move away from simple case of just one omitted variable, determining sign (and magnitude) of bias will be a lot harder
 - Let β be vector of coefficients on k included variables
 - Let γ be vector of coefficient on l excluded variables
 - Let \mathbf{X} be matrix of observations of included variables
 - Let \mathbf{Z} be matrix of observations of excluded variables

$$\hat{\beta} = \beta + \frac{E[\mathbf{X}'\mathbf{Z}]}{E[\mathbf{X}'\mathbf{X}]} \gamma$$

OVB – General Form, Intuition

$$\hat{\beta} = \beta + \frac{E[X'Z]}{E[X'X]} \gamma$$

Vector of regression
coefficients

Vector of partial effects of
excluded variables

- Same idea as before, but more complicated
- Frankly, this can be a real mess!

[See Gormley and Matsa (2014) for example with just two included and two excluded variables]

Eliminating Omitted Variable Bias

- How we try to get rid of this bias will depend on the type of omitted variable

- **Observable** omitted variable

- **Unobservable** omitted variable



**How can we deal with an
observable omitted variable?**

Observable omitted variables

- This is easy! Just add them as controls
 - E.g. if the omitted variable, z , in my simple case was ‘leverage,’ then add leverage to regression
- A functional form misspecification is a special case of an observable omitted variable

Let’s now talk about this...

Functional form misspecification

- Assume true model is...

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_2^2 + u$$

- But, we omit squared term, x_2^2
 - Just like any OVB, bias on $(\beta_0, \beta_1, \beta_2)$ will depend on β_3 and correlations among (x_1, x_2, x_2^2)
 - You get same type of problem if have incorrect functional form for y [e.g. *it should be $\ln(y)$ not y*]
- In some sense, this is minor problem... Why?

Tests for correction functional form

- You could add additional squared and cubed terms and look to see whether they make a difference and/or have non-zero coefficients
- This isn't as easy when the possible models are not nested...

Non-nested functional form issues...


- Two non-nested examples are:

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + u$$

versus

$$y = \beta_0 + \beta_1 \ln(x_1) + \beta_2 \ln(x_2) + u$$

Let's use this example and see how we can try to figure out which is right



$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + u$$

versus

$$y = \beta_0 + \beta_1 x_1 + \beta_2 z + u$$

Davidson-MacKinnon Test [Part 1]

- To test which is correct, you can try this...
 - Take fitted values, \hat{y} , from 1st model and add them as a control in 2nd model

$$y = \beta_0 + \beta_1 \ln(x_1) + \beta_2 \ln(x_2) + \theta_1 \hat{y} + u$$

- Look at t -stat on θ_1 ; if significant rejects 2nd model!
- Then, do reverse, and look at t -stat on θ_1 in

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \theta_1 \hat{y} + u$$

where \hat{y} is predicted value from 2nd model... if significant then 1st model is also rejected ☹

Davidson-MacKinnon Test *[Part 2]*

- Number of weaknesses to this test...
 - A clear winner may not emerge
 - Both might be rejected
 - Both might be accepted [If this happens, you can use the R^2 to choose which model is a better fit]
 - And, rejecting one model does **NOT** imply that the other model is correct ☹

Bottom line advice on functional form

- Practically speaking, you hope that changes in functional form won't effect coefficients on key variables very much...
 - But, if it does... You need to think hard about why this is and what the correct form should be
 - The prior test might help with that...

Eliminating Omitted Variable Bias

- How we try to get rid of this bias will depend on the type of omitted variable
 - **Observable** omitted variable
 - **Unobservable** omitted variable



Unobservable are much harder to deal with, but one possibility is to find a proxy variable

Unobserved omitted variables

- Again, consider earlier estimation

$$\ln(wage) = \beta_0 + \beta_1 educ + \beta_2 ability + u$$

- **Problem:** we don't observe & can't measure *ability*
- What can we do? **Ans. =** Find a proxy variable that is correlated with the unobserved variable, E.g. IQ

Proxy variables [Part 1]

- Consider the following model...

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3^* + u$$

where x_3^* is unobserved, but we have proxy x_3

- Then, suppose $x_3^* = \delta_0 + \delta_1 x_3 + v$
 - v is error associated with proxy's imperfect representation of unobservable x_3
 - Intercept just accounts for different scales
[e.g. ability has different average value than IQ]

Proxy variables [Part 2]

- If we are only interested in β_1 or β_2 , we can just replace x_3^* with x_3 and estimate

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + u$$

- But, for this to give us consistent estimates of β_1 and β_2 , we need to make some assumptions

#1 – We've got the right model, and

#2 – Other variables don't explain unobserved variable after we've accounted for our proxy

Proxy variables – *Assumptions*

#1 – $E(u | x_1, x_2, x_3^*) = 0$; i.e. we have the right model and x_3 would be irrelevant if we could control for x_1, x_2, x_3^* , such that $E(u | x_3) = 0$

□ This is a common assumption; not controversial

#2 – $E(v | x_1, x_2, x_3) = 0$; i.e. x_3 is a good proxy for x_3^* such that after controlling for x_3, x_3^* doesn't depend on x_1 or x_2

□ I.e. $E(x_3^* | x_1, x_2, x_3) = E(x_3^* | x_3)$

Why the proxy works...

- Recall true model: $y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3^* + u$
- Now plug-in for x_3^* , using $x_3^* = \delta_0 + \delta_1 x_3 + v$

$$y = \underbrace{(\beta_0 + \beta_3 \delta_0)}_{\alpha_0} + \beta_1 x_1 + \beta_2 x_2 + \underbrace{(\beta_3 \delta_1)}_{\alpha_1} x_3 + \underbrace{(u + \beta_3 v)}_e$$

- Prior assumptions ensure that $E(e | x_1, x_2, x_3) = 0$
such that the estimates of $(\alpha_0, \beta_1, \beta_2, \alpha_1)$ are consistent
- **Note:** β_0 and β_3 are not identified

Proxy assumptions are key [Part 1]

- Suppose assumption #2 is wrong such that

$$x_3^* = \delta_0 + \delta_1 x_3 + \underbrace{\gamma_1 x_1 + \gamma_2 x_2 + w}_v$$

where $E(w | x_1, x_2, x_3) = 0$

- If above is true, $E(v | x_1, x_2, x_3) \neq 0$, and if you substitute into model of y , you'd get...

Proxy assumptions are key [Part 2]

- Plugging in for x_3^* , you'd get

$$y = \alpha_0 + \alpha_1 x_1 + \alpha_2 x_2 + \alpha_3 x_3 + e$$

where $\alpha_0 = \beta_0 + \beta_3 \delta_0$

$$\alpha_1 = \beta_1 + \beta_3 \gamma_1$$

$$\alpha_2 = \beta_2 + \beta_3 \gamma_2$$

$$\alpha_3 = \beta_3 \delta_1$$

E.g. α_1 captures effect of x_1 on y , β_1 , but also its correlation with unobserved variable

- We'd get consistent estimates of $(\alpha_0, \alpha_1, \alpha_2, \alpha_3)$
But that isn't what we want!

Proxy variables – *Example #1*

- Consider earlier wage estimation

$$\ln(\text{wage}) = \beta_0 + \beta_1 \text{educ} + \beta_2 \text{ability} + u$$

- If use IQ as proxy for unobserved *ability*, what assumption must we make? Is it plausible?
 - **Answer:** We assume $E(\text{ability} | \text{educ}, \text{IQ}) = E(\text{ability} | \text{IQ})$, i.e. average ability does not change with education after accounting for IQ... Could be questionable assumption!

Proxy variables – *Example #2*

- Consider Q -theory of investment

$$\textit{investment} = \beta_0 + \beta_1 Q + u$$

- Can we estimate β_1 using a firm's market-to-book ratio (MTB) as proxy for Q ? Why or why not?
 - **Answer:** Even if we believe this is the correct model (Assumption #1) or that Q only depends on MTB (Assumption #2), e.g. $Q = \delta_0 + \delta_1 \text{MTB}$, we are still not getting estimate of β_1 ... see next slide for the math

Proxy variables – *Example #2 [Part 2]*

- Even if assumptions held, we'd only be getting consistent estimates of

$$investment = \alpha_0 + \alpha_1 MTB + e$$

where $\alpha_0 = \beta_0 + \beta_1 \delta_0$

$$\alpha_1 = \beta_1 \delta_1$$

- While we can't get β_1 , is there something we can get if we make assumptions about sign of δ_1 ?
- **Answer:** Yes, the sign of β_1

Proxy variables – Summary

- If the coefficient on the unobserved variable isn't what we are interested in, then a proxy for it can be used to identify and remove OVB from the other parameters
 - Proxy can also be used to determine sign of coefficient on unobserved variable

Random Coefficient Model

- So far, we've assumed that the effect of x on y (i.e. β) was the same for all observations
 - In reality, this is unlikely true; model might look more like $y_i = \alpha_i + \beta_i x_i + u_i$, where

$$\alpha_i = \alpha + c_i$$

$$\beta_i = \beta + d_i$$

$$E(c_i) = E(d_i) = 0$$

I.e. each observation's relationship between x and y is slightly different

- α is the average intercept and β is what we call the “average partial effect” (APE)

Random Coefficient Model [Part 2]

- Regression would seem to be incorrectly specified, but if willing to make assumptions, we can identify the APE

- Plug in for α_i and β_i

$$y_i = \alpha + \beta x_i + (c_i + d_i x_i + u_i)$$

- Identification requires

$$E(c_i + d_i x_i + u_i | x) = 0$$

What does this imply?

← If like, can think of the unobserved differential intercept and slopes as omitted variable

Random Coefficient Model [Part 3]

- This amounts to requiring

$$E(c_i | x) = E(c_i) = 0 \Rightarrow E(\alpha_i | x) = E(\alpha_i)$$

$$E(d_i | x) = E(d_i) = 0 \Rightarrow E(\beta_i | x) = E(\beta_i)$$

- We must assume that the individual slopes and intercepts are mean independent (i.e. uncorrelated with the value of x) in order to estimate the APE
 - I.e. knowing x , doesn't help us predict the individual's partial effect

Random Coefficient Model [Part 4]

- Implications of APE
 - Be careful interpreting coefficients when you are implicitly arguing elsewhere in paper that effect of x varies across observations
 - **Keep in mind the assumption this requires**
 - **And, describe results using something like...**
“we find that, on average, an increase in x causes a β change in y ”

Three main ways this will be violated

- Omitted variable bias
- Measurement error bias
- Simultaneity bias

Measurement error (ME) bias

- Estimation will have measurement error whenever we measure the variable of interest imprecisely
 - Ex. #1: Altman-z-score is noisy measure of default risk
 - Ex. #2: Avg. tax rate is noisy measure of marg. tax rate
- **Such measurement error can cause bias, and the bias can be quite complicated**

Measurement error *vs.* proxies

- Measurement error is similar to proxy variable, but very different conceptually
 - Proxy is used for something that is entirely unobservable or measureable (e.g. ability)
 - With measurement error, the variable we don't observe is well-defined and can be quantified... it's just that our measure of it contains error

ME of Dep. Variable [Part 1]

- Usually not a problem (in terms of bias); just causes our standard errors to be larger. E.g.,...
- Let $y^* = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + u$
- But, we measure y^* with error $e = y - y^*$
- Because we only observe y , we estimate

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + (u + e)$$



Note: we always assume $E(e)=0$; this is innocuous because if untrue, it only affects the bias on the constant

ME of Dep. Variable [Part 2]

- As long as $E(e | x) = 0$, the OLS estimates are consistent and unbiased
 - I.e. as long as the measurement error of y is uncorrelated with the x 's, we're okay
 - Only issue is that we get larger standard errors when e and u are uncorrelated [*which is what we typically assume*] because $\text{Var}(u+e) > \text{Var}(u)$

What are some common examples of ME?

ME of Dep. Variable *[Part 3]*

- Some common examples
 - **Market leverage** – typically use book value of debt because market value hard to observe
 - **Firm value** – again, hard to observe market value of debt, so we use book value
 - **CEO compensation** – value of options are approximated using Black-Scholes

Is assuming e and x are uncorrelated plausible?

ME of Dep. Variable *[Part 4]*

- **Answer** = Maybe... maybe not
 - Ex. – Firm leverage is measured with error; hard to observe market value of debt, so we use book value
 - But, the measurement error is likely to be larger when firm's are in distress... Market value of debt falls; book value doesn't
 - This error could be correlated with x's if it includes things like profitability (i.e. ME larger for low profit firms)
 - **This type of ME will cause inconsistent estimates**

ME of Independent Variable [Part 1]

- Let's assume the model is $y = \beta_0 + \beta_1 x^* + u$
- But, we observe x^* with error, $e = x - x^*$
 - We assume that $E(y | x^*, x) = E(y | x^*)$ [i.e. x doesn't affect y after controlling for x^* ; this is standard and uncontroversial because it is just stating that we've written the correct model]
- **What are some examples in CF?**

ME of Independent Variable *[Part 2]*

- There are lots of examples!
 - Average Q measures marginal Q with error
 - Altman- ζ score measures default prob. with error

Will this measurement error cause bias?

ME of Independent Variable *[Part 2]*

- Answer depends crucially on what we assume about the measurement error, e
- Literature focuses on two extreme assumptions
 - #1 – Measurement error, e , is uncorrelated with the observed measure, x
 - #2 – Measurement error, e , is uncorrelated with the unobserved measure, x^*

Assumption #1: e uncorrelated with x

- Substituting x^* with what we actually observe, $x^* = x - e$, into true model, we have

$$y = \beta_0 + \beta_1 x + u - \beta_1 e$$

- Is there a bias?

- **Answer** = No. x is uncorrelated with e by assumption, and x is uncorrelated with u by earlier assumptions

- What happens to our standard errors?

- **Answer** = They get larger; error variance is now $\sigma_u^2 + \beta_1^2 \sigma_e^2$

Assumption #2: e uncorrelated with x^*

- We are still estimating $y = \beta_0 + \beta_1 x + u - \beta_1 e$, but now, x is correlated with e
 - e uncorrelated with x^* guarantees e is correlated with x ; $\text{cov}(x, e) = E(xe) = E(x^* e) + E(e^2) = \sigma_e^2$
 - I.e. an independent variable will be correlated with the error... we will get **biased** estimates!
- This is what people call the **Classical Error-in-Variables (CEV)** assumption

CEV with 1 variable = attenuation bias

- If work out math, one can show that the estimate of β_1 , $\hat{\beta}_1$, in prior example (which had just one independent variable) is...

$$p \lim(\hat{\beta}_1) = \beta_1 \left(\frac{\sigma_{x^*}^2}{\sigma_{x^*}^2 + \sigma_e^2} \right) \longleftarrow \text{This scaling factor is always between 0 and 1}$$

- The estimate is always biased towards zero; i.e. it is an **attenuation bias**
 - And, if variance of error, σ_e^2 , is small, then attenuation bias won't be that bad

Measurement error... not so bad?

- Under current setup, measurement error doesn't seem so bad...
 - If error uncorrelated with observed x , no bias
 - If error uncorrelated with unobserved x^* , we get an attenuation bias... so at least the sign on our coefficient of interest is still correct
- Why is this misleading?

Nope, measurement error is bad news

- Truth is, measurement error is probably correlated a bit with both the observed x and unobserved x^*
 - I.e... some attenuation bias is likely
- **Moreover**, even in CEV case, if there is more than one independent variable, the bias gets horribly complicated...

ME with more than one variable

- If estimating $y = \beta_0 + \beta_1 x_1 + \dots + \beta_k x_k + u$, and just one of the x 's is mismeasured, then...
 - **ALL** the β 's will be biased if the mismeasured variable is correlated with any other x
[which presumably is true since it was included!]
 - Sign and magnitude of biases will depend on all the correlations between x 's; **i.e. big mess!**
 - See Gormley and Matsa (2014) math for *AvgE* estimator to see how bad this can be

ME example

- Fazzari, Hubbard, and Petersen (1988) is classic example of a paper with ME problem
 - Regresses investment on Tobin's Q (it's measure of investment opportunities) and cash
 - Finds positive coefficient on cash; argues there must be financial constraints present
 - But Q is noisy measure; all coefficients are biased!
- Erickson and Whited (2000) argues the pos. coeff. disappears if you correct the ME

Three main ways this will be violated

- Omitted variable bias
- Measurement error bias
- Simultaneity bias

Simultaneity bias

- This will occur whenever any of the supposedly independent variables (i.e. the x 's) can be affected by changes in the y variable; E.g.

$$y = \beta_0 + \beta_1 x + u$$

$$x = \delta_0 + \delta_1 y + v$$

- I.e. changes in x affect y , and changes in y affect x ; this is the simplest case of reverse causality
- An estimate of $y = \beta_0 + \beta_1 x + u$ will be biased...

Simultaneity bias continued...

- To see why estimating $y = \beta_0 + \beta_1 x + u$ won't reveal the true β_1 , solve for x

$$x = \delta_0 + \delta_1 y + v$$

$$x = \delta_0 + \delta_1 (\beta_0 + \beta_1 x + u) + v$$

$$x = \left(\frac{\delta_0 + \delta_1 \beta_0}{1 - \delta_1 \beta_1} \right) + \left(\frac{v}{1 - \delta_1 \beta_1} \right) + \left(\frac{\delta_1}{1 - \delta_1 \beta_1} \right) u$$

- Easy to see that x is correlated with u ! I.e. bias!

Simultaneity bias in other regressors

- Prior example is case of reverse causality; the variable of interest is also affected by y
- But, if y affects any x , there will be a bias; E.g.

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + u$$

$$x_2 = \gamma_0 + \gamma_1 y + w$$

- Easy to show that x_2 is correlated with u ; and there will be a bias on all coefficients
- This is why people use lagged x 's

“Endogeneity” problem – *Tangent*

- In my opinion, the prior example is what it means to have an “endogeneity” problem or an “endogenous” variable
 - But, as I mentioned earlier, there is a lot of misuse of the word “endogeneity” in finance... So, it might be better just saying “simultaneity bias”

Simultaneity Bias – Summary

- If your x might also be affected by the y (i.e. reverse causality), you won't be able to make causal inferences using OLS
 - Instrumental variables or natural experiments will be helpful with this problem
- Also can't get causal estimates with OLS if controls are affected by the y

Summary of Today [Part 1]

- We need conditional mean independence (CMI), to make causal statements
- CMI is violated whenever an independent variable, x , is correlated with the error, u
- Three main ways this can be violated
 - Omitted variable bias
 - Measurement error bias
 - Simultaneity bias

Summary of Today *[Part 2]*

- The biases can be very complex
 - If more than one omitted variable, or omitted variable is correlated with more than one regressor, sign of bias hard to determine
 - Measurement error of an independent variable can (and likely does) bias all coefficients in ways that are hard to determine
 - Simultaneity bias can also be complicated

Summary of Today *[Part 3]*

- To deal with these problems, there are some tools we can use
 - E.g. Proxy variables [discussed today]
 - We will talk about other tools later, e.g.
 - Instrumental variables
 - Natural experiments
 - Regression discontinuity

In 1st Half of Next Class

- Before getting to these other tools, will first discuss panel data & unobserved heterogeneity
 - Using fixed effects to deal with unobserved variables
 - What are the benefits? [There are many!]
 - What are the costs? [There are some...]
 - Fixed effects versus first differences
 - When can FE be used?

In 2nd Half of Next Class

- Instrumental variables
 - What are the necessary assumptions? [E.g. what is the exclusion restriction?]
 - Is there are way we can test whether our instruments are okay?